

# Local Ensemble Transform Kalman Filter

Brian Hunt

11 June 2013

# Review of Notation

- Forecast model: a known function  $M$  on a vector space of model states.
- Truth: an unknown sequence  $\{x_n\}$  of model states to be estimated.
- Observations: a sequence  $\{y_n^o\}$  of vectors in observation space.
- Forward operator: a known function  $H_n$  from model space to observation space.
- Observation error:  $\varepsilon_n = y_n^o - H_n(x_n)$ .
- Background (before data): superscript  $b$ .
- Analysis (after data): superscript  $a$ .

# Kalman Filter Review

- The Kalman Filter tracks a mean  $x_n^{b,a}$  and covariance  $P_n^{b,a}$  representing a Gaussian distribution of model states.
- The analysis step assumes Gaussian observation errors and computes (exactly if  $H_n$  is linear)  $x_n^a$  and  $P_n^a$  from  $x_n^b$ ,  $P_n^b$ , and the observation error covariance  $R_n$ .
- The forecast step sets  $x_{n+1}^b = M(x_n^a)$ .
- If  $M$  is nonlinear, there is no “right” way to propagate the covariance from  $P_n^a$  to  $P_{n+1}^b$ .

# Practical Difficulties

- If  $x$  is  $m$ -dimensional, then  $P$  is an  $m$ -by- $m$  matrix
- If  $m$  is large, the computational cost of storing and manipulating  $P$  may be prohibitive.
- Linearizing the model around the forecast trajectory (as required by the Extended Kalman Filter) can be expensive.
- Ensemble Kalman Filters make a low-rank approximation to  $P$  and avoid explicit linearization of the forecast model.

# Ensemble Kalman Filters

- Introduced by Geir Evensen [1994].
- Use the sample mean and covariance of an ensemble of forecasts as the Kalman filter background.
- Form an ensemble of states that matches the Kalman Filter analysis mean and covariance and use it to initialize the next ensemble forecast.
- Assumption: the model is approximately linear within the range of the ensemble.

# Ensemble Analysis

- Subscripts now index ensemble members and not time.
- Input: background ensemble  $x_1^b, \dots, x_k^b$  with sample mean  $\bar{x}^b$  and covariance  $P^b$ ; observation information  $y^o, H, R$ .
- Output: analysis ensemble  $x_1^a, \dots, x_k^a$  with sample mean  $\bar{x}^a$  and covariance  $P^a$  determined as in the Kalman Filter.

# Kalman Filter Equations

- Recall the Kalman filter analysis equations:

$$K = P^b H^T [H P^b H^T + R]^{-1}$$

$$\bar{x}^a = \bar{x}^b + K(y^o - H\bar{x}^b)$$

$$P^a = (I - KH)P^b$$

- Note:  $\bar{x}^a - \bar{x}^b$  is called the **analysis increment**.
- What should we do if  $H$  is nonlinear?
- How do we determine the analysis ensemble?

# Nonlinear Forward Operator $H$

- Write  $P^b = X^b(X^b)^T$  where  $X^b$  is the matrix whose  $i$ th column is  $(x_i^b - \bar{x}^b)/\sqrt{k-1}$ .
- Let  $Y^b = HX^b$  if  $H$  is linear.; then:

$$K = X^b(Y^b)^T [Y^b(Y^b)^T + R]^{-1}$$

$$\bar{x}^a = \bar{x}^b + K(y^o - H\bar{x}^b)$$

$$P^a = (X^b - KY^b)(X^b)^T$$

- For nonlinear  $H$ , let  $y_i^b = H(x_i^b)$ , let  $\bar{y}^b$  be the mean of the  $y_i^b$ , and let  $Y^b$  be a matrix whose  $i$ th column is  $(y_i^b - \bar{y}^b)/\sqrt{k-1}$ .

# Perturbed Observations

- There are many possible analysis ensembles whose mean and covariance are  $\bar{x}^a$  and  $P^a$ .
- One can apply the Kalman Filter update to each ensemble member, perturbing the observed values differently for each:

$$x_i^a = x_i^b + K(y^o + \varepsilon_i - y_i^b)$$

where  $\varepsilon_i \sim N(0, R)$  [Burgers et al. 1998, Houtekamer & Mitchell 1998].

- Then if  $H$  is linear, one can show that the **expected** mean and covariance of the analysis ensemble are  $\bar{x}^a$  and  $P^a$ .

# Deterministic Approaches

- **Square root filters** track a matrix “square root” of the model state covariance matrix; in our case  $X^b$  for which  $P^b = X^b(X^b)^T$ .
- We seek a matrix  $X^a$  of (scaled) analysis ensemble perturbations such that  $P^a = X^a(X^a)^T$ .
- Add  $\bar{x}^a$  to the (scaled) columns of  $X^a$  to get the analysis ensemble.
- Various approaches to determining  $X^a$  from  $X^b$ : EAKF [Anderson 2001], ETKF [Bishop et al. 2001, Wang et al. 2004], EnSRF [Whitaker & Hamill 2002]; see also [Tippett et al. 2003].

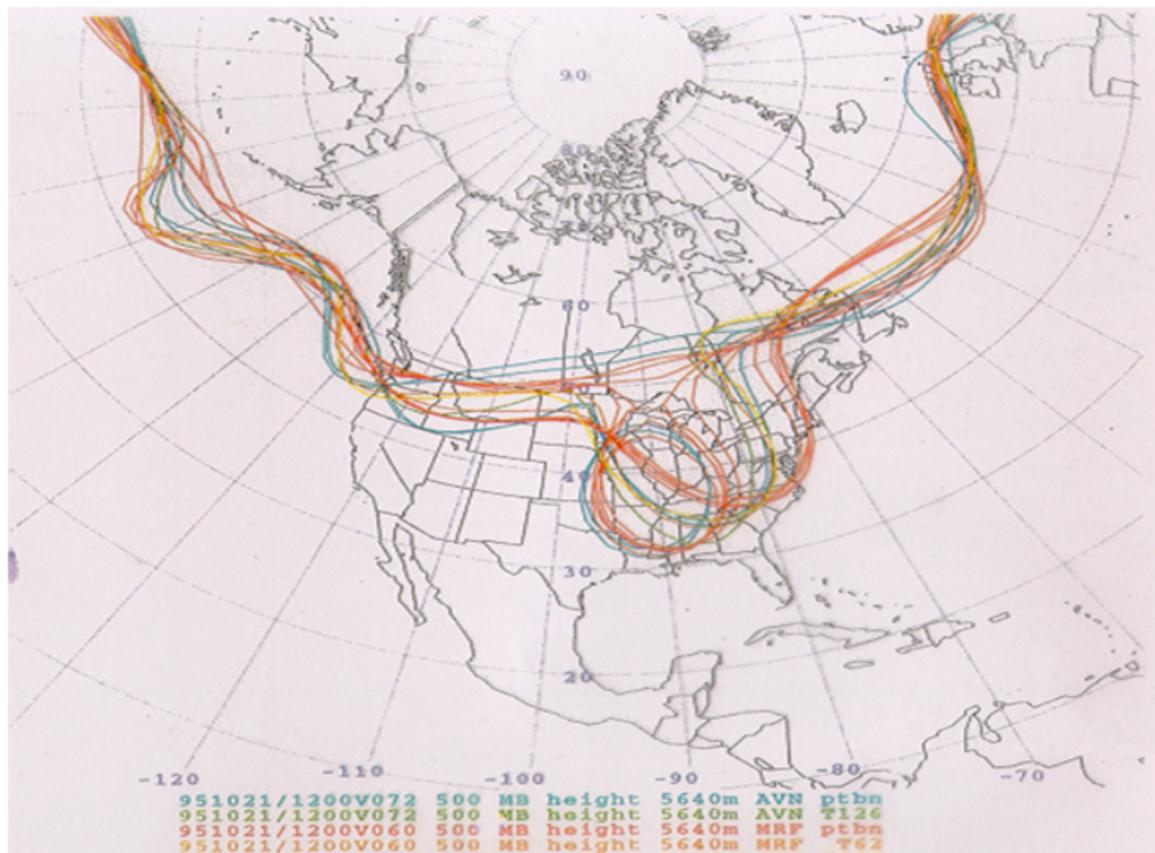
# Reduced Rank: Pros and Cons

- The background covariance allows analysis increments only in the space spanned by the ensemble perturbations.
- Indeed, in all the approaches on the last two slides, each analysis ensemble member can be written  $x_i^a = \bar{x}^b + X^b w_i$ .
- If the ensemble is too small, the filter may fail.
- If an ensemble of reasonable size is successful, analysis computations can be done in the **ensemble space** spanned by the background ensemble perturbations.

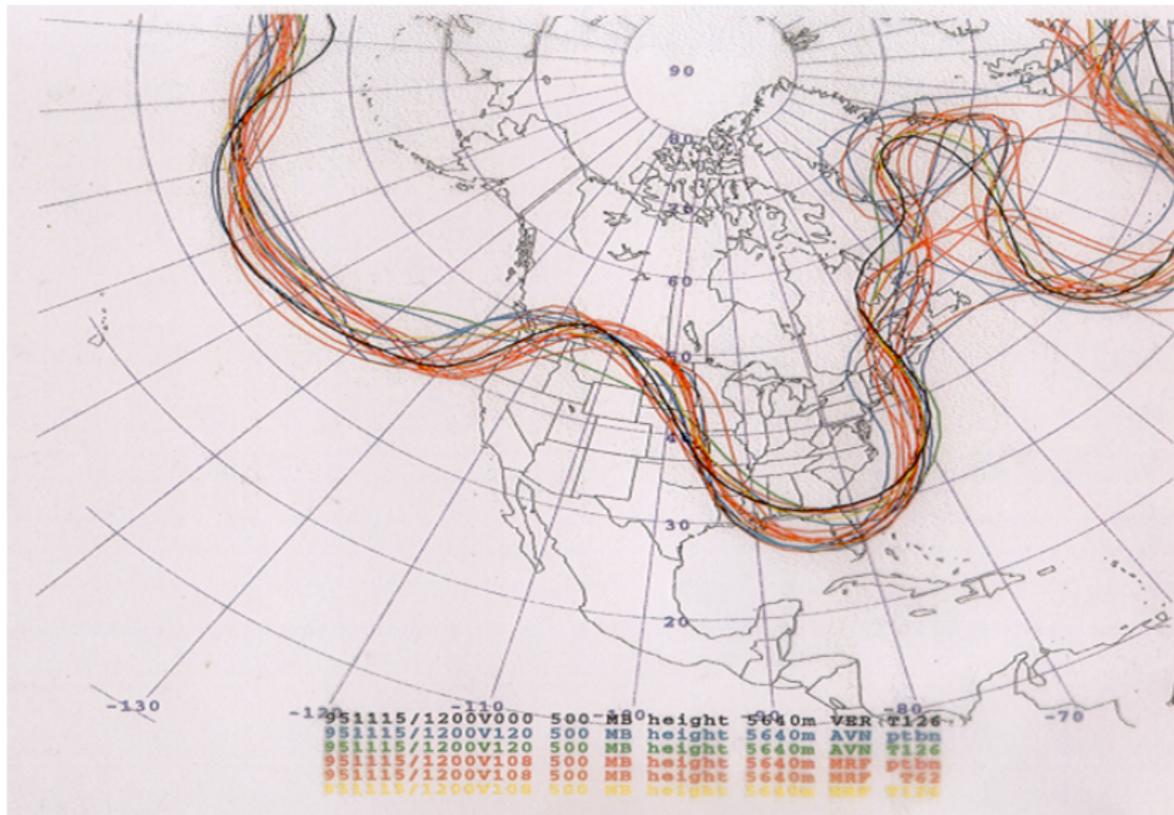
# Underlying Assumptions

- A linear combination of background ensemble states is a plausible state.
- The truth lies reasonably close to the ensemble space.
- More precisely, most of the uncertainty in the background state lies in the ensemble space.
- In particular, forecast uncertainty lies in a relatively low-dimensional space.

# A 2.5 Day Ensemble Forecast



# A 4.5 Day Ensemble Forecast



# Spatial Localization

- For a spatially extended system, it may not be feasible to forecast an ensemble large enough to span the possible global states.
- If long-distance correlations are weak, the ensemble covariance will represent spurious long-distance correlations.
- **Localization** reduces the effect of spurious correlations and allows analysis increments from a higher dimensional space.

# Localization Approaches

- A standard approach is to localize the background covariance  $P^b$  by multiplying each element by a factor between 1 (for variables at the same grid point) and 0 (for variables farther apart than a chosen **localization distance**).
- Our approach [LEKF, Ott et al. 2004; LETKF, Hunt et al. 2007] is do a separate analysis for each grid point using only observations from within a local region that we choose.
- These analyses can be done in parallel.

# Inflation does “Time Localization”

- Many filters use **multiplicative covariance inflation**, which at each step multiplies the forecast (background) covariance by  $\rho$  for an ad hoc parameter  $\rho > 1$ .
- This is equivalent to multiplying the term corresponding to observations from  $t$  analysis cycles ago in the least squares cost function by a weight  $\rho^{-t}$ .
- Covariance inflation compensates for effects of model error, model nonlinearity, etc.

# LETKF Formalism

- LETKF stands for Local Ensemble Transform Kalman Filter; equivalent to LEKF but formulated more like ETKF.
- Recall  $P^b = X^b(X^b)^T$  where  $X^b$  is a matrix whose  $i$ th column is  $(x_i^b - \bar{x}^b)/\sqrt{k-1}$ .
- Consider model states  $x = \bar{x}^b + X^b w$  where  $w$  is a  $k$ -dimensional vector.
- If  $w$  has mean  $0$  and covariance  $I$ , then  $x$  has mean  $\bar{x}^b$  and covariance  $P^b$ .

# Analysis in Ensemble Space

- Let  $y_i^b = H(x_i^b)$ , let  $\bar{y}^b$  be their mean, and form the matrix  $Y^b$  of perturbations like  $X^b$ .
- Make the linear approximation:

$$H(\bar{x}^b + X^b w) \approx \bar{y}^b + Y^b w$$

- Minimize the cost function:

$$J(w) = \rho^{-1} w^T w + (\bar{y}^b + Y^b w - y^o)^T R^{-1} (\bar{y}^b + Y^b w - y^o)$$

- Compare to:

$$J(x) = (x - \bar{x}^b)^T (\rho P^b)^{-1} (x - \bar{x}^b) + (H(x) - y^o)^T R^{-1} (H(x) - y^o)$$

# Analysis in Ensemble Space

- The analysis mean  $\bar{w}^a$  and covariance  $A$  are:

$$A = [I + (Y^b)^T R^{-1} Y^b]^{-1}$$
$$\bar{w}^a = A(Y^b)R^{-1}(y^o - \bar{y}^b)$$

- Then  $\bar{x}^a = \bar{x}^b + X^b \bar{w}^a$  and  $P^a = X^b A (X^b)^T$ .
- Notice that the analysis equations in  $w$  coordinates depend only on the background ensemble  $\{y_i^b\}$  in observation space and the observation data  $y^o$  and  $R$ .
- The matrix that is inverted to find  $A$  is  $k$ -by- $k$  and has no small eigenvalues.

# Asynchronous Observations

- With this formulation, the filter easily becomes “4D” (properly takes into account temporal information) when observations are asynchronous.
- When mapping the background ensemble members  $x_i^b$  into observation space, use the background state at the appropriate time for each observation.
- Assumption: a linear combination of ensemble trajectories is an approximate model trajectory.

# Choice of Analysis Ensemble

- To form the analysis ensemble vectors  $w_i^a$ , add to  $\bar{w}^a$  the columns of the symmetric square root  $W^a = [(k - 1)A]^{1/2}$ .
- The analysis ensemble  $x_i^a = \bar{x}^b + X^b w_i^a$  then has the correct mean and covariance.
- Other choices are possible, but this choice minimizes the change (in  $w$ , or in  $x$  with  $P^b$ -norm) between the background and analysis ensembles.
- Recall that we determine a different  $w_i^a$  for each grid point, using only observations from a local region near that grid point.

# Weight Interpolation

- If the localization distance is large relative to the grid spacing, overlapping local regions make nearby analyses consistent.
- For a high-resolution model, the overlap causes a lot of computational redundancy.
- Solution: Compute the weights  $w_i^a$  on a coarse grid of analysis points and interpolate them to the other grid points.
- Can improve balance [Yang et al. 2009].

# Tapering Observation Influence

- At a given grid point, our local analysis (as described so far) uses nearby observations at “full” influence, while distant observations have zero influence.
- The influence can be tapered more smoothly from “1” to 0 by multiplying each diagonal entry  $\sigma_j^{-2}$  in  $R^{-1}$  by a factor  $0 \leq \alpha_j \leq 1$  that depends on the distance from the corresponding observation location to the analysis grid point.
- If  $R$  is not diagonal, replace  $R^{-1}$  by  $DR^{-1}D$ , where  $D$  is diagonal with entries  $\sqrt{\alpha_j}$ .

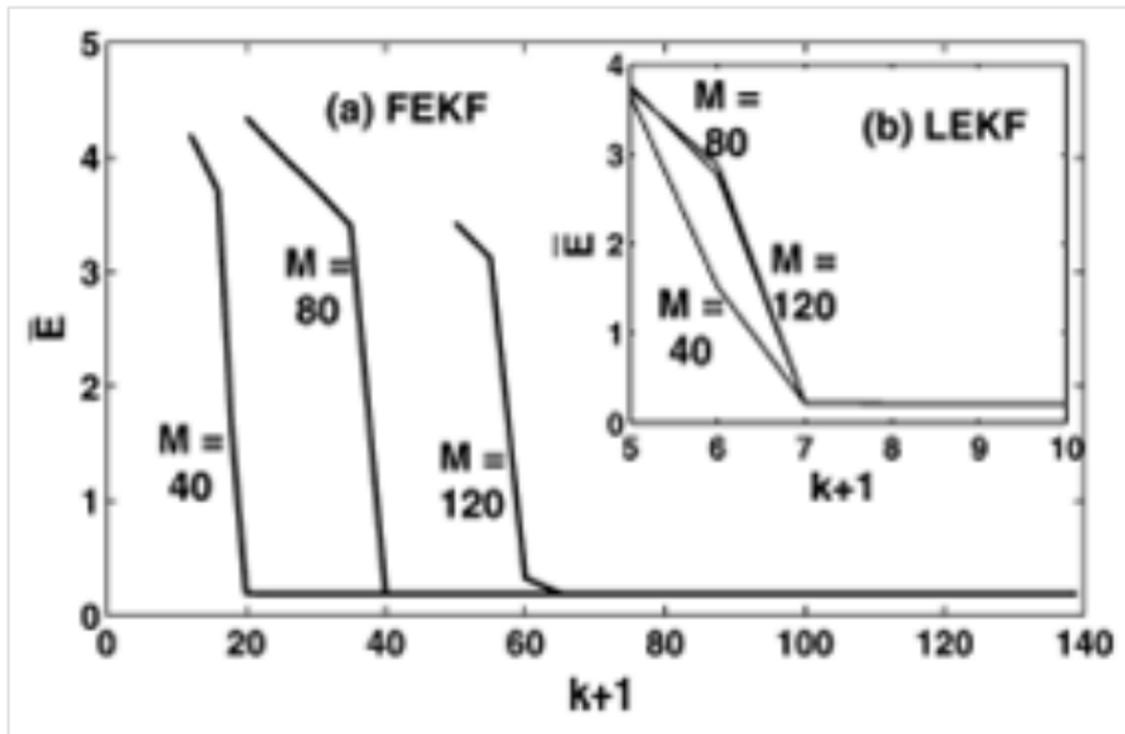
# Toy Model Results

- Perfect model test with Lorenz-96 system of coupled ODEs on a circular lattice:

$$dx_m/dt = (x_{m+1} - x_{m-2})x_{m-1}x_m + 8.$$

- We used a model trajectory as the truth and add noise to simulate observations.
- We compared our LEKF, using only observations from within 6 grid points, with a global ensemble Kalman filter (FEKF) [Ott et al. 2004].

# Global vs. Local Filter



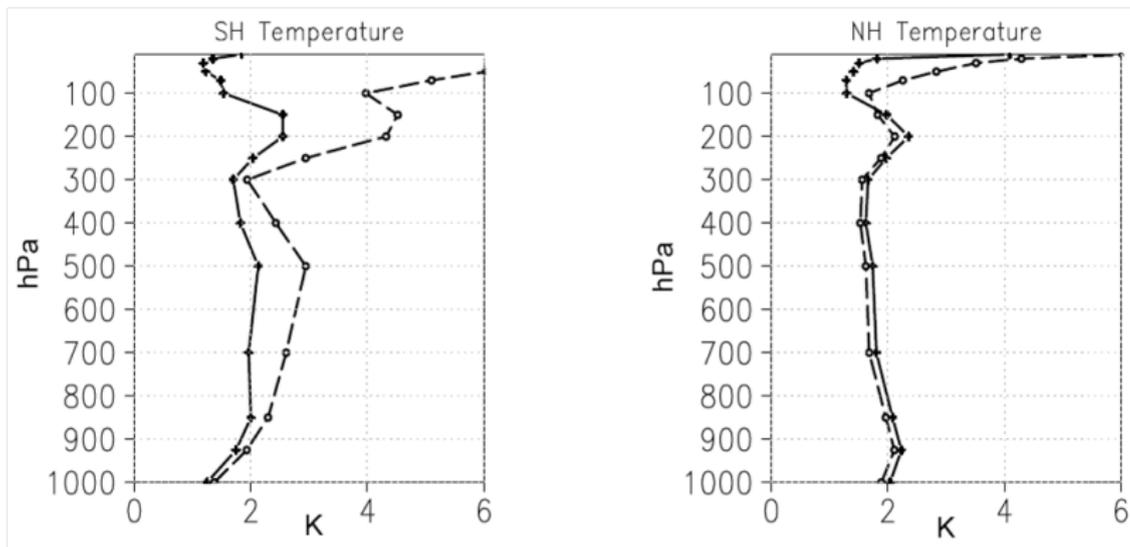
$M$  = model size,  $k+1$  = ensemble size,  $E$  = error

# Comparison to NCEP (2004) 3D-Var

- We ran the U.S. National Centers for Environmental Prediction global forecast model (GSM) at reduced T62 resolution (about 500,000 grid points).
- We compared our LETKF analyses with analyses from NCEP's 3D-Var system (SSI) using actual Winter 2004 non-satellite observations (about 300,000 per 6 hours).
- We used a 60-member ensemble, 800 km radius local region, spatially varying inflation  $\rho \sim 1.25$  [Szunyogh et al. 2008].

# 48-hour Forecast Error

(compared to radiosonde observations)



+ = our LETKF

o = NWS 3D-Var

# Computational Speed

- At NCEP, less than 10 minutes of every 6 hour cycle is used for data assimilation.
- Our implementation took about 15 minutes on a 40-processor Linux cluster.
- The computation time is approximately:
  - **linear** in the number of observations;
  - **linear** in the number of model grid points;
  - **quadratic** in the number of ensemble members.

# Extensions

- Parameter/Bias Estimation: treat parameters of  $M$  or  $H$  as state variables with time derivative zero (Baek et al. 2006, 2009; Cornick et al. 2009; Fertig et al. 2009).
- Nongaussian Filter: minimize nonquadratic cost function numerically in ensemble space (Harlim & Hunt 2007), as in MLEF (Zupanski 2006).

# Conclusions

- LETKF is a relatively simple, efficient, and flexible framework for reduced-rank data assimilation that works in practice.
- The method is largely model-independent.
- It scales well to large systems.
- Full citations for references to our group's work are on the publications page at:  
<http://www.weatherchaos.umd.edu>
- Motto: Forecast globally, assimilate locally.